# Testing the power of an experiment to measure predator control and habitat complexity impacts on farmland bird abundance 

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#### Abstract

In this study I assess the statistical power to detect a significantly greater increase in bird population size on treatment farms than on control farms given that there is a substantial treatment effect. Computer simulations of bird populations on New Zealand sheep/beef farms were used to generate significant changes in bird abundance from (a) controlling predation by introduced small mammals, (b) habitat structural complexity, and (c) an interaction of both. A simplified computer model of bird population dynamics was developed that predicted a birth pulse of $357 \%$ when predators were controlled and $110 \%$ if not, and a target of detecting the experimental elevation of bird abundance at a statistically significant level ( $P<0.05$ ) in $75 \%$ of all attempts was set. If at least four farm pairs (treatment vs non-treatment) are monitored, this is feasible for 15 of 23 species common on farmlands for which sampling error of abundance estimation was below $\sim 40 \%$. A second virtual experiment measured the power of tests of whether habitat complexity and predation in combination led to added increases in bird abundance. It showed that a $75 \%$ detection of elevated benefits of predation control in complex habitats could only be achieved if at least 48 farms were monitored, and then only for species for which abundance could be estimated with $<10 \%$ error. Researchers are advised to invest in increased withinsite monitoring to achieve a reasonable precision in bird abundance estimation before increasing the number of replicates.


Keywords: abundance estimation; agricultural biodiversity; New Zealand; power analysis; simulation

## Introduction

Enhancement of biodiversity on farmed landscapes in New Zealand undoubtedly depends on building more habitat variety and structural complexity into existing agro-ecosystems (Price 1993; Meurk \& Swaffield 2000; Perley et al. 2001; Blackwell et al. 2008; Moller et al. 2008a, b; Weller et al. 2008). However, this strategy will not be sufficient in itself if the introduced mammalian predators that now inhabit New Zealand's farming landscapes suppress populations below the levels where habitat would limit their populations. Many farmers may resist planting of woody vegetation to create ecological refuges from agricultural disturbance because it potentially reduces the area of their farms that produces food, fibre and profits. The Agriculture Research Group on Sustainability (ARGOS) therefore proposed an experiment to test whether predator control was a necessary additional intervention for habitat restoration to increase bird abundance. ARGOS's overall goal is to provide incentives for the restoration of bird diversity and abundance in the production landscapes of New Zealand by demonstrating the comparative efficacy of predation control, habitat enhancement and a mixture of both. Feral cats (Felis catus), possums (Trichosurus vulpecula), feral ferrets (Mustela furo), stoats (M. erminea), weasels (M. nivalis), Norway rats (Rattus norvegicus), ship rats ( $R$. rattus), mice (Mus musculus) and hedgehogs (Erinaceus europaeus) are all potentially threatening valued indigenous and introduced species on farms (Towns \& Ballantine 1993; Cowan \& Tyndale-Biscoe 1997; Perley et al. 2001; King 2005; Moller et al. 2008b). The greatest danger for birds is predation of eggs and nestlings, or even the incubating parent, during the breeding season (Martin 1993, 1995).

The power to measure experimental effects is partly determined by the number of replicates of each treatment, the duration of the experiment, and uncertainty in the monitoring method used to measure the response variables (in this case the abundance of birds). Large-scale field experiments are very expensive and logistically challenging to complete with adequate replication and duration. Most management of pest control operations in Australia and New Zealand does not have enough replication or monitoring of response variables in nontreatment plots to be scientifically interpretable (Reddiex et al. 2006; Clayton \& Cowan 2009). A survey of field experiments published in top international ecology journals found that most had too few replicates of treatment plots, many did not even have a non-treatment comparison, and very few lasted for the generation time of the species being monitored for response to the experimental perturbation (Moller \& Raffaelli 1998; Raffaelli \& Moller 2000). Very few researchers conducted formal power analyses before mounting these experiments, and some that did a power analysis then ignored the result and performed the experiment even though the analysis predicted that the result would be virtually uninterpretable.

The lack of prior planning and formal power analysis of a proposed experimental design may also partly reflect a lack of knowledge of input parameters and measures of their variance before the experiment is performed. Furthermore, many proposed experiments are so complex that power analyses are difficult to conduct. One potential way to circumvent these problems is to set up 'virtual experiments' in computer simulations and systematically vary assumptions and design parameters. This paper describes a power analysis based on the simulation of two 'press perturbation' (Bender et al. 1984) virtual experiments that would test whether predator control will
significantly increase the abundance of birds in New Zealand farmland. The simulations focused strongly on uncertainty in bird abundance estimates and how this interacts with the number of replicates of the experiment to affect the ability to detect the experimental effect. My primary research goal was to assess the statistical power to (a) detect a statistically significant increase in bird abundance in $75 \%$ of all attempts, and (b) determine whether any effects of predator control were statistically significantly different in simple compared with complex habitats in $75 \%$ of attempts, depending on the number of experimental replicates and magnitude of the error in bird abundance estimates.

The farms envisaged for the experiments are a group of 36 sheep and beef farms on the South Island of New Zealand, including a subset used for intensive year-round bird abundance monitoring. They are part of the farms that are being studied by the ARGOS project (Moller et al. 2005, Weller 2009). The focal farms considered for the experiment could be broadly categorised as having 'low complexity' or 'high complexity' habitat types. 'Low complexity' farms are composed mostly of open paddocks, are of limited topographical variety and have little or no native vegetation; any introduced woody vegetation is mostly in the form of shelterbelts. 'High complexity' farms have areas of native forest, pine plantations or heterogeneous patches, and include gullies or otherwise hard-to-access terrain that provide refuges for birds and predators alike.

## Simulated study system

## Experimental structure

Two variants of an experiment based on the same fundamental set-up were envisaged. Over a period of several years, the nesting success of selected bird species would be monitored on replicated sets of farms. After establishing a 'baseline' value during the first breeding season (spring and summer of the first year), as many predators as is feasible would be removed by trapping from one-half of the selected farms while constantly monitoring predator numbers on all farms. This 'predator press' treatment would be maintained for 2 years. ABefore-After-Control-Impact design (BACI; Stewart-Oaten \& Murdoch 1986) would be used to monitor changes in bird abundance on both sets of farms, permitting a comparison of conditions before and after the experimental manipulation in one set while also providing constant comparison with a non-manipulated control set.

Experiment $A$ would focus on establishing just the effect of predation control for a given complexity of habitat. 'Treatment' and 'control' farms with comparable habitats would be grouped into replicated pairs, and the change in bird population densities compared within each pair.

Experiment $B$ would expand this to use groups of four farms, consisting of (1) a 'treatment'/'control' pair with low habitat diversity and a simple habitat structure, and (2) a pair with varied and complex habitats. This would allow the investigation of interaction effects of habitat quality and changes in predation pressure on bird populations.

## Simulation procedure

A computer simulation was created to model the expected dynamics of the monitored bird populations. The object of this simulation was to test whether a change in bird density caused by the predator control on the 'treatment' farms $(A)$, or
the difference in density change between habitat complexity levels $(B)$, would be detectable by standard statistical methods; i.e. to test the statistical power of the proposed experimental set-up. The simulation program was written in Microsoft QuickBasic 7.1 (Microsoft) and provided a very simplified numerical model of a bird population over the proposed duration of the study.

The simulation created the change in a population density over 3 years. At the beginning of each year, the population went through a brief period of increase that corresponded to recruitment following a 'birth pulse' during the breeding season (following Caughley 1967). The pulse encompasses the laying of eggs, incubation, and hatching, and was set at $10 \%$ of total year duration. During the rest of the year, population size declined to minimum adult density just before the next breeding season, representing dispersal and deaths in a population at equilibrium. To ensure easy modifiability and simulation speed, only the maximum (just after the birth pulse) and minimum (just before the birth pulse) values for each year were calculated by using the population model, and the data points between were interpolated using a standardised exponential curve.

The magnitude of the standard birth pulse was simulated by predicting nesting success under different predation pressures. Since the experimentally variable part of predation pressure in this scenario consisted of predation of eggs and nestlings only, the population size at any point in the model represented the number of adult individuals. Therefore all simulated predation took place during the birth pulse and determined its peak height.

The simulations were run in two parallel streams. For experiment $A$, each stream represented a 'treatment' and 'control' (no predator control) farm as a matched pair of farms for each replication. For experiment $B$, each of the two treatment types consisted of pairs of farms with 'high habitat complexity' and 'low habitat complexity', making up farm quartets. In all scenarios I simulated a year of population fluctuation (the 'Before' phase of the BACI) and then imposed predator control on the 'treatment' farm(s) for two successive years. From the time that predator control went into effect, the estimated losses of eggs and chicks due to predation were removed on 'treatment' farms (i.e. the birth pulse increased), while the annual rate of population decrease due to dispersal and deaths from other causes was maintained at 'Before' level. This resulted in a cumulative increase of population size on treatment farms over the following years, for as long as predator control was continued (Fig. 1).

## Simulation parameters

Two types of parameters entered into the models (Table 1): a model parameter set describing the modelled population (birth pulse strengths, and variability in pulse strengths and in starting population size), and the experimental parameters describing the set-up of the experiment and the monitoring process (number of replicates, simulation years, bird count samples taken per year, and introduced sampling error).

## Estimating birth pulses and starting population sizes

Population model parameters estimated characteristics of bird populations and were obtained from published literature. Using data on clutch sizes and number of nestings in Heather \& Robertson (2000), an average maximum reproduction rate was calculated across 23 species likely to be encountered


Figure 1. Simulation of bird population size over the course of 3 years, on the 'treatment' (top) and 'control' (below) farm. Equilibrium populations increase in a short 'birth pulse' during nesting season and then decline back to the starting value over the rest of the year. The vertical line marks the onset of predator control on the 'treatment' farm, resulting in a reduction of nest predation and an increase in birth pulse strength. The taking of virtual bird count samples to represent sampling errors in the field is shown in year 3 (circles). Samples are generated by adding a random error (between $-10 \%$ and $10 \%$ in this example) to the calculated population size.

Table 1. Population model parameters and experimental parameters used in the simulations, with sources as applicable.

| Simulation $A$ <br> Population parameters <br> Parameter | Value | Sources | Experimental parameters <br> Parameter |
| :--- | :--- | :--- | :--- | :--- |
| Standard birth pulse strength <br> (reduced by nest mortality) | $+109 \%$ | Heather \& Robertson (2000) | Value |

on the monitored farms in sufficient numbers for abundance estimation (Table 2). This value was halved for simulation purposes to account for a population of breeding pairs. The list of species was based on the results of previous ARGOS bird surveys on the intended properties (Blackwell et al. 2005; Green et al. 2005).

These reproduction rates predicted an average birth pulse of $+375 \%$ of the basal adult population just before breeding began if no eggs or chicks were lost (Table 2). A standard birth pulse strength (i.e. under nest predation) was then calculated as $+109 \%$ of the pre-breeding adult population size, this being the proportion of individuals successfully hatched after factoring in an average $71 \%$ of nest losses reported in the literature (various sources, see Table 1). Birth pulse strength when predators were controlled was then calculated as $+308 \%$ by removing the estimated $75 \%$ of these total nest losses that were attributable to nest predation (leaving a residual nest mortality of $18 \%$ ) (Martin 1993). A maximum random variation of $50 \%$ of starting population sizes within a farm pair was inferred from the ARGOS results and a maximum random variation of $50 \%$ of annual birth pulse from Heather \& Robertson (2000).

For simulation $B$, the starting population size in 'high complexity' habitats was assumed to be $140 \%$ of that in 'low complexity' habitats, due to greater availability of nesting and feeding sites (based on a 2-year study carried out on ARGOS farms; Weller 2009). The identical 'before' (and continuous 'control') birth pulse value was assigned to both complexity levels, as the source data that were used to calculate the reproduction rates (Heather \& Robertson 2000) did not allow segregation between habitat types, and
no information is available at this time to adjust this value in the case of New Zealand farmland birds. Since the purpose of this simulation was to assess detection power in the case of a hypothetical differential effect of predator removal on nesting success being present, populations on 'high complexity' 'treatment' farms were assigned an increased reaction to the release from predation pressure, based on the assumption that benefits from predation release would be all the more realised in complex habitats (Begon et al. 1996; Whittingham \& Evans 2004). Series were produced for differences of $10 \%, 20 \%, 30 \%$ and $40 \%$. Half of each value was subtracted from the 'low complexity' 'treatment' birth pulse strength and half added to the 'high complexity' one since the reproduction rates averaged from literature were sourced from a variety of habitat types (Table 1); e.g. for a difference of $20 \%$, the treatment pulse strength of $308 \%$ would be modified to $298 \%$ for the 'low complexity' site and $318 \%$ for the 'high complexity' site.

Simulation $A$ runs were performed with 2,3 and 4 farm pair replicates, which would allow implementation on the ARGOS core farms. Simulation $B$ proved to require a substantially larger amount of replication and was performed with 4,8 and 12 replicates of farm quartets.

## Incorporating uncertainty from monitoring bird abundance

To simulate uncertainty from field monitoring of bird abundance, sampling error in abundance estimation was simulated rather than assuming perfect knowledge of the number of birds present. Simulated bird count samples were taken twice a year, at the beginning and end of the post-birthpulse period. A variable error (percentage CV) was applied to the current state value of the model to simulate sampling errors

Table 2. Population model parameters used to calculate magnitude of reproduction rates, for all species where abundance estimates could be produced in the ARGOS surveys (Blackwell et al. 2005; Green et al. 2005). Species where abundance could be estimated with an error below $40 \%$ (percentage CV) are marked with an asterisk *. Data are sourced from Heather \& Robertson (2000). The reproduction value is the annual average percentage population increase per parent bird based on number of broods and eggs.

| Common name | Scientific name | Avg. broods | Avg. eggs |  | Avg. reproduction \% |
| :---: | :---: | :---: | :---: | :---: | :---: |
| Bellbird* | Anthornis melanura | 2 | 3 |  | 300 |
| Blackbird* | Turdus merula | 2.5 | 3 |  | 375 |
| Chaffinch* | Fringilla coelebs | 2 | 4 |  | 400 |
| Dunnock | Prunella modularis | 2.5 | 4 |  | 500 |
| Fantail* | Rhipidura fuliginosa | 2.5 | 3 |  | 375 |
| Feral pigeon | Columba livia | 2.5 | 2 |  | 250 |
| Goldfinch* | Carduelis carduelis | 1.5 | 4 |  | 300 |
| Greenfinch* | Carduelis chloris | 2 | 5 |  | 500 |
| Grey warbler* | Gerygone igata | 2 | 4 |  | 400 |
| Australasian harrier* | Circus approximans | 1 | 3 |  | 150 |
| House sparrow* | Passer domesticus | 3 | 4 |  | 600 |
| Australian magpie* | Gymnorhina tibicen | 1 | 3 |  | 150 |
| Mallard duck | Anas platyrhynchos | 1.5 | 9 |  | 675 |
| Paradise shelduck | Tadoma variegata | 1 | 8 |  | 400 |
| Pied oystercatcher | Haematopus finschi | 1 | 2 |  | 100 |
| Red poll* | Carduelis flammea | 2 | 4 |  | 400 |
| Silvereye | Zosterops lateralis | 2 | 3 |  | 300 |
| Skylark* | Alauda arvensis | 2.5 | 3 |  | 375 |
| Song thrush* | Turdus philomelos | 2.5 | 3 |  | 375 |
| Spur-winged plover* | Vanellus miles | 2.5 | 4 |  | 500 |
| Starling | Sturnus vulgaris | 1.5 | 4 |  | 300 |
| Welcome swallow | Hirundo tahitica | 2.5 | 4 |  | 500 |
| Yellowhammer | Emberiza citrinella | 2 | 4 |  | 400 |
|  |  |  |  | Average | 375 |

incurred in the field, and this adjusted value was recorded (Fig. 1). Runs used random values from a range of simulated error maxima between $0 \%$ and $\pm 200 \%$. In accordance with actual bird monitoring practice on the sheep/beef farms intended for the experiment, where data from several surveys would be pooled for distance modelling of an annual abundance estimate, the two samples per year were averaged into single annual values before analysis.

## Testing power to detect virtual experimental effects

The samples taken from a simulation run of a unit of farms were analysed with SPSS 10.0 (SPSS Inc. 1999) ( $A$ ) and GenStat 9.1 (Lawes Agricultural Trust 2006) (B). ‘Before’ values were subtracted from 'After' values of the equivalent part of the year, and the differences used in modelling. For simulation $A$, a repeated-measures general linear model (identity link function, normal errors) of the form year*treatmentType was used, where treatmentType was either 'treatment' or 'control'. Since interest lay in the ability to distinguish population size changes on 'treatment' farms after implementation of predator control from those on 'control' farms over the same period, the frequency of significant effects (at the 0.05 level) of the interaction factor year.treatmentType were noted. Thus, the abundance difference was modelled as $\alpha_{f}+\beta_{\mathrm{i}} * t$, where $\alpha_{f}$ was the intercept depending on starting population size of farm $f, t$ was the year, $i$ was either 'treatment' or 'control', and the tested null hypothesis was $\beta_{\text {treatment }}=\beta_{\text {control }}$. Similarly, for simulation $B$, the model had the form year*treatmentType*complexity, where complexity was either 'high' or 'low', and significant effects of the three-way interaction year.treatmentType. complexity were tested. The simulation run for a unit of farms was repeated 100 times, and the percentage of replicates that showed a significant interaction effect was recorded. For simulation $A$, this process was carried out for each of 13 different introduced sampling error sizes, and the resulting series was
repeated for 2, 3 and 4 pairs of farm replicates. Simulation $B$ used five different error levels within the range shown to yield interpretable results for simulation $A(0 \%-40 \%)$, and 4,8 or 12 farm quartet replicates (Table 1 ).

## Simulation results

As expected, in both simulations statistical power to detect the effect of the predator control treatment declined with increasing bird count uncertainty (Figs $2 \& 3$ ). A higher number of farm pair replicates strongly increased the power to detect the effect, especially if the sampling error of abundance estimation was also low. Over the parameter space explored by the simulation, the sampling error at first had a much greater effect on experimental power than did the level of replication; below an error level of $\sim 100 \%$, additional replications rapidly increased in their effect on power in simulation $A$ (Fig. 2).

However, in the case of simulation $A$, even if estimation techniques were very certain (e.g. sampling error was $<10 \%$ ), having just two replicates will only detect the experimental effect in around $40 \%$ of cases (Fig. 2). Setting a target minimum percentage of detections at $75 \%$, the simulation shows that at least four farm pairs must be monitored provided that the maximum sampling error of the density estimates does not exceed $\sim 40 \%$ (Fig. 2).

Simulation $B$ showed that power to detect effects of habitat complexity on the strength of benefit from predator control was much lower for a given number of replicates than for the detection of the predator control effect alone. Four replicates of farm quartets yielded a detection probability of only $30 \%$ even assuming perfect accuracy in bird population size estimation and the largest hypothetical habitat-related difference in birth pulse strengths (40\%) (Fig. 3). At the sampling error maximum for simulation $A$ of $40 \%$, detection probability in $B$ was below


Figure 2. Power to detect the effects of predator control underan increasing size of sampling error for bird abundance estimates. $X$ axis displays maximum simulated sampling error, $y$-axis displays percentage of 100 simulation runs using thaterror where a significant (0.05) effect of predator control was detected. Run series with 2,3 and 4 replicate farm pairs are represented by individual data series (smoothed splines). Maximum sampling error to achieve a power of $75 \%$ using four farm pair replicates is marked by a dashed line.


Figure 3. Power to detect the interaction effects of predator control and habitat complexity under an increasing size of sampling error for bird abundance estimates. $X$-axis displays maximum simulated sampling error, $y$-axis displays percentage of 100 simulation runs using that error where a significant (0.05) interaction effect was detected. Run series with $10 \%, 20 \%, 30 \%$, and $40 \%$ difference in 'treatment' birth pulse between 'high' and 'low complexity' farms are represented by individual data series. Panels represent results for 4,8 and 12 farm quartet replicates. Maximum sampling error to achieve a power of $75 \%$ using 12 farm quartet replicates is marked by a dashed line.
$50 \%$ for every parameter combination. The power threshold of $75 \%$ could be achieved only with 12 replicates ( 48 farms), but only if a birth pulse strength difference of $40 \%$ and a sampling error not larger than $10 \%$ were assumed (Fig. 3).

## Discussion

The very simplified nature of my simulation needs to be taken into account when considering its reliability. Using two deterministic population targets per year, the program makes no allowance for more complex aspects of dynamics that are frequently present in bird populations, such as density dependence (Begon et al. 1996) or the possibility of renesting if one clutch or nest is lost (Thompson et al. 2001), nor for any kind of seasonal population fluctuations other than a generalised diminishing of numbers from the breeding season onwards. Nevertheless, the actual impact of inaccuracies in density distribution across a year should be small, as neither the sampling nor the analysis process make any assumptions about the temporal correlation between bird count samples within a year. Also, my main interest lay in the relative difference in abundance between reference and experimental farms, not the absolute values. Errors in assumptions will apply equally to both groups and are therefore less likely to have affected my predictions of necessary replication.

A potentially more serious bias could result from miscalculation of the magnitude of the effect of predation (or its lack). I used averaged values from the literature. Repeating this analysis for a specific species, with more accurate population parameters should they be available, should give better information for individual cases and might result in modified conclusions for measuring the response of a particular species to predator control, while this study was primarily concerned with examining the general feasibility of the proposed experiments.

The simulation results obtained for experiment $A$, assessing the detectability of the effects of predator removal irrespective of habitat type, indicate that for an experiment aiming to restore any of the species potentially present in New Zealand farmland, using generalised assumptions, a satisfactory degree of statistical power could be achieved if density estimates are sufficiently accurate. The best available estimates of the actual sampling errors when monitoring bird numbers on these farms can be derived from three recent annual monitoring studies carried out on the ARGOS farms (ARGOS, unpubl. data). In these surveys, the abundance of 15 of the 23 most commonly encountered bird species could be estimated with less than $40 \%$ sampling error (using line transect distance sampling; Buckland et al. 2001, 2004) (Table 2). If four farm unit replicates were used, this level of estimation error would be within the desirable power of detection range. The $75 \%$ power threshold chosen is appropriate, given the expense of undertaking a press-perturbation experiment. While this means that the putative experimental test of the benefits of predator control could not be reliable for about a third of these common species, power could be increased by working with ecologically sensible species groupings sharing similar detectability and/or life history traits. Collectively, my results (Fig. 2) suggest that a considerable proportion of the investment in any predator experiment should go into increased frequency and therefore accuracy in the bird monitoring itself. A high number of replicates will not suffice to raise the power of the experiment to interpretable levels if a reasonable accuracy of estimates cannot be achieved. This is also procedurally more sensible, in that a well-designed, accurate monitoring programme can be extended to new areas more easily than a widespread, inaccurate one can be upgraded to better estimates.

Testing the interaction effect of predator removal and habitat quality, as proposed for experiment $B$, appears to be far less feasible. Carrying out sustained predator control on 24 farms, and monitoring bird populations on twice as many, for
the duration of 3 years would represent a high investment of effort even if there was a high chance of successfully measuring the sought effect. However, the parameters chosen for the simulation already represent almost a best-case scenario for the detection of an interaction effect. Assuming that there is a higher base population in 'high complexity' farms even under predation pressure is supported by the results of surveys on representative farms (Weller 2009), but may not represent the general case. Repeating the simulation run with the highest detection power ( $40 \%$ treatment birth pulse difference, 12 replicates, no estimation error) without a bonus applied to 'high complexity' starting populations yielded a power level of only $15 \%$ instead of $77 \%$, showing that in the absence of this supposition the interaction effect would be at least fivefold harder to detect. As noted above, the error present in estimates of bird abundance is also unlikely to consistently be as low as $10 \%$ even using comparatively accurate distance sampling methods (Weller 2009; ARGOS, unpubl. results). Additionally, the set of 36 farms where these estimates were taken would not be large enough to carry out the experiment at the needed rate of replication.

Most importantly, the strength of the hypothetical effect difference between complexity levels being unknown, there would be a high probability of not detecting the sought interaction if the difference in relative population increase was smaller than the assumed $40 \%$. A negative effect size, i.e. a relatively reduced increase of the birth pulse on 'high complexity' farms, for example caused by the population size already being closer to the habitat's carrying capacity, would similarly be much harder to detect.

## Conclusion

Power analyses based on simulated bird population dynamics showed that experimental investigations of the effect of predation on the breeding success of farmland birds would require four replicated pairs of farms to obtain $75 \%$ certainty of results, which would be feasible using the bird surveying techniques already tested on the intended experimental farms. An expanded variant of the experiment that would investigate the interaction effect of predator control and habitat quality on bird breeding success is not feasible due to expectations of low power to detect the effect even at high replication. Simulations to estimate experimental power of large-scale field manipulations have considerable scope for improving wildlife science and gaining more reliable knowledge.

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